

## COMP8601: Advanced Topics in Theoretical Computer Science

### Lecture 7: Johnson-Lindenstrauss Lemma: Dimension Reduction

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Date: 8 Oct 2013

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*These lecture notes are supplementary materials for the lectures. They are by no means substitutes for attending lectures or replacement for your own notes!*

## 1 Dimension Reduction in Euclidean Space

Consider  $n$  vectors in Euclidean space of some large dimension. These  $n$  vectors reside in an  $n$  dimensional subspace. By rotation, we can assume that  $n$  vectors lie in  $\mathbb{R}^n$ . On the other hand, it is easy to see that  $n$  mutually orthogonal unit vectors cannot reside in a space with dimension less than  $n$ .

Moreover, it is not possible to have three mutually almost orthogonal vectors placed in 2 dimensions.

**Definition 1.1** We say two unit vectors  $u$  and  $v$  are  $\epsilon$ -orthogonal to one another if their dot product satisfies  $|u \cdot v| \leq \epsilon$ .

One might think that  $n$  mutually almost orthogonal vectors require  $n$  dimensions. Hence, it might come as a surprise that  $n$  vectors that are mutually  $\epsilon$ -orthogonal can be placed in a Euclidean space with  $O(\frac{\log n}{\epsilon^2})$  dimensions.

Observe that for any three points, if the three distances between them are given, then the three angles are fixed. Given  $n - 1$  vectors, the vectors together with the origin form a set of  $n$  points. In fact, given any  $n$  points in Euclidean space (in  $n - 1$  dimensions), the Johnson-Lindenstrauss Lemma states that the  $n$  points can be placed in  $O(\frac{\log n}{\epsilon^2})$  dimensions such that distances are preserved with multiplicative error  $\epsilon$ , for any  $0 < \epsilon < 1$ .

**Theorem 1.2 (Johnson-Lindenstrauss Lemma [JL84])** Suppose  $U$  is a set of  $n$  points in Euclidean space  $\mathbb{R}^n$ . Then, for any  $0 < \epsilon < 1$ , there is a mapping  $f : U \rightarrow \mathbb{R}^T$ , where  $T = O(\frac{\log n}{\epsilon^2})$ , such that for all  $x, y \in U$ ,

$$(1 - \epsilon)||x - y||^2 < ||f(x) - f(y)||^2 < (1 + \epsilon)||x - y||^2.$$

**Remark 1.3** 1. Since for small  $\epsilon$ ,  $(1 + \epsilon)^2 = 1 + \Theta(\epsilon)$  and  $(1 - \epsilon)^2 = 1 - \Theta(\epsilon)$ , it follows that the squared of the distances are preserved iff the distances themselves are.

2. Note that  $||x - y||$  is a norm between 2 vectors in Euclidean space  $\mathbb{R}^n$  and  $||f(x) - f(y)||$  is one between 2 vectors in  $\mathbb{R}^T$ . Be careful that,  $||x - f(x)||$  is not well-defined.

**Corollary 1.4 (Almost Orthogonal Vectors)** Suppose  $u_1, u_2, \dots, u_n$  are mutually orthogonal unit vectors in  $\mathbb{R}^n$ . Then, for any  $0 < \epsilon < 1$ , there exists a mapping  $f : U \rightarrow \mathbb{R}^T$ , where  $T = O(\frac{\log n}{\epsilon^2})$  such that  $|\frac{f(u_i)}{||f(u_i)||} \cdot \frac{f(u_j)}{||f(u_j)||}| \leq \epsilon$ .

**Proof:** We apply Johnson-Lindenstrauss' Lemma with error  $\frac{\epsilon}{8}$  to the set  $U$  of vectors  $u_1, u_2, \dots, u_n$  together with the origin to obtain  $f : U \rightarrow \mathbb{R}^T$ , where  $T = O(\frac{\log n}{\epsilon^2})$ .

Hence, it follows that for all  $i$ ,  $1 - \frac{\epsilon}{8} \leq \|f(u_i)\|^2 \leq 1 + \frac{\epsilon}{8}$ .

Moreover, for  $i \neq j$ ,  $(1 - \frac{\epsilon}{8})\|u_i - u_j\|^2 < \|f(u_i) - f(u_j)\|^2 < (1 + \frac{\epsilon}{8})\|u_i - u_j\|^2$ .

Observe that  $\|u_i - u_j\|^2 = 2$  and  $\|f(u_i) - f(u_j)\|^2 = \|f(u_i)\|^2 + \|f(u_j)\|^2 - 2f(u_i) \cdot f(u_j)$ .

So, from  $(1 - \frac{\epsilon}{8})\|u_i - u_j\|^2 < \|f(u_i) - f(u_j)\|^2$ , we conclude  $f(u_i) \cdot f(u_j) \leq \frac{\epsilon}{4}$ .

On the other hand, from  $\|f(u_i) - f(u_j)\|^2 < (1 + \frac{\epsilon}{8})\|u_i - u_j\|^2$ , we have  $f(u_i) \cdot f(u_j) \geq -\frac{\epsilon}{4}$ .

Hence, we have  $|f(u_i) \cdot f(u_j)| \leq \frac{\epsilon}{4}$ . However, observe that  $f(u_i)$  and  $f(u_j)$  might not be unit vectors. We know that  $\|f(u_i)\| \cdot \|f(u_j)\| \geq (1 - \frac{\epsilon}{8})^2 \geq \frac{1}{4}$ . Therefore, we have  $|\frac{f(u_i)}{\|f(u_i)\|} \cdot \frac{f(u_j)}{\|f(u_j)\|}| \leq \epsilon$ .  
■

## 2 Random Projection

Several proofs [DG03, Ach03] of the theorem are based on random projection. The construction can be derandomized [EIO02], but the argument is quite involved.

For point  $x$ , suppose  $f(x) := (f_i(x))_{i \in [T]}$ . Then,  $\|f(x) - f(y)\|^2 = \sum_{i \in [T]} |f_i(x) - f_i(y)|^2$ .

We have learned that the sum of independent random variables concentrate around its mean. Hence, the goal is to design a random mapping  $f_i : U \rightarrow \mathbb{R}$  such that  $E[|f_i(x) - f_i(y)|^2] = \frac{1}{T} \cdot \|x - y\|^2$ , in which case we have  $E[\|f(x) - f(y)\|^2] = \|x - y\|^2$ .

Note that  $f_i$  takes a vector and returns a number. Observe that Euclidean space is equipped with dot product. Note that dot product with a unit vector gives the magnitude of the projection on the unit vector. Hence, we can take a random vector  $r$  in space  $\mathbb{R}^n$ , and let  $f_i$  have the form  $f_i(x) := r \cdot x$ .

Suppose we fix two points  $x$  and  $y$ . Since dot product is linear, we have  $f_i(x) - f_i(y) = f_i(x - y)$ . Hence, we consider  $v := x - y = (v_0, v_1, \dots, v_{n-1})$ , and let  $\nu := \|v\| = \sqrt{\sum_i v_i^2}$ . Recall the goal is to define  $f_i$ , and hence find a random vector  $r$  such that  $E[(r \cdot v)^2] = \frac{1}{T} \cdot \|v\|^2 = \frac{\nu^2}{T}$ .

**Using Random Bits to Define a Random Projection.** The following idea of using random bits is due to Achlioptas [Ach03]. For each  $j \in [n]$ , suppose  $\gamma_j \in \{-1, +1\}$  is a uniform random bit such that  $\gamma$ 's are independent. Define the random vector  $r := \frac{1}{\sqrt{T}}(\gamma_0, \gamma_1, \dots, \gamma_{n-1})$ . Hence,  $f_i(v) = \frac{1}{\sqrt{T}} \sum_j \gamma_j v_j$ .

Check that  $E[(f_i(v))^2] = \frac{1}{T} \sum_j v_j^2 = \frac{\nu^2}{T}$ . Hence, we have found the required random mapping  $f_i : \mathbb{R}^n \rightarrow \mathbb{R}^T$ .

**Remark 2.1** Observe that the mapping  $f : \mathbb{R}^n \rightarrow \mathbb{R}^T$  is linear.

## 3 Proof of Johnson-Lindenstrauss Lemma

We define  $X_i := f_i(v)^2 = \frac{1}{T}(\sum_j \gamma_j v_j)^2$ , and let  $Y := \sum_i X_i$ . Recall  $E[X_i] = \frac{\nu^2}{T}$  and  $E[Y] = \nu^2$ . Then, the desirable event can be expressed as:

$$Pr[(1 - \epsilon)\|x - y\|^2 < \|f(x) - f(y)\|^2 < (1 + \epsilon)\|x - y\|^2] = Pr[|Y - E[Y]| < \epsilon E[Y]].$$

The goal is to first find a  $T$  large enough such that the failing probability  $Pr[|Y - E[Y]| \geq \epsilon E[Y]]$  is at most  $\frac{1}{n^2}$ . Since there are  $\binom{n}{2}$  such pairs of points, using union bound, we can show that with probability at least  $\frac{1}{2}$ , the distances of all pairs of points are preserved.

We again use the method of moment generating function.

### 3.1 JL as a Measure Concentration Result

Using the method of moment generating function described in previous classes, the failure probability in question is at most the sum of the following two probabilities.

1.  $Pr[Y \leq (1 - \epsilon)\nu^2] \leq \exp(-t(1 - \epsilon)\nu^2) \cdot \prod_i E[\exp(tX_i)]$ , for all  $t < 0$ .
2.  $Pr[Y \geq (1 + \epsilon)\nu^2] \leq \exp(-t(1 + \epsilon)\nu^2) \cdot \prod_i E[\exp(tX_i)]$ , for all  $t > 0$ .

We next derive an upper bound for  $E[e^{tX_i}]$ .

## 4 Upper Bound for $E[e^{tX_i}]$

For notational convenience, we drop the subscript  $i$ , and write  $X := \frac{1}{T}(\sum_j \gamma_j v_j)^2$ , where  $\nu^2 = \sum_j v_j^2$ , where  $\gamma_j \in \{-1, 1\}$  are uniform and independent. Hence, we have

$$E[e^{tX}] = E[\exp(\frac{t}{T}(\sum_j v_j^2 + \sum_{i \neq j} \gamma_i \gamma_j v_i v_j))].$$

Although the  $\gamma_j$ 's are independent, the cross-terms  $\gamma_i \gamma_j$ 's are not. In particular,  $\gamma_i \gamma_j$  and  $\gamma_{i'} \gamma_{j'}$  are not independent if  $i = i'$  or  $j = j'$ .

We compare  $X$  with another variable  $\hat{X}$ , which we can analyze.

### 4.1 Normal Distribution

Suppose  $g$  is a random variable having standard normal distribution  $N(0, 1)$ , with mean 0 and variance 1. In particular, it has the following probability density function:

$$\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, \text{ for } x \in \mathbb{R}.$$

Suppose  $\gamma$  is a  $\{-1, 1\}$  is a random variable that takes value  $-1$  or  $1$ , each with probability  $\frac{1}{2}$ . Then, the random variables  $g$  and  $\gamma$  have some common properties.

**Fact 4.1** Suppose  $\gamma$  is a uniform  $\{-1, 1\}$ -random variable and  $g$  is a random variable with normal distribution  $N(0, 1)$ .

1.  $E[\gamma] = E[g] = 0$ .
2.  $E[\gamma^2] = E[g^2] = 1$ .

For higher moments we have,

1. For odd  $n \geq 3$ ,  $E[\gamma^n] = E[g^n] = 0$ .

2. For even  $n \geq 4$ ,  $1 = E[\gamma^n] \leq E[g^n]$ .

Normal distributions have the following important property.

**Fact 4.2** Suppose  $g_i$ 's are independent random variables, each having standard normal distribution  $N(0, 1)$ . Define  $Z := \sum_j g_j v_j$ , where  $v_j$ 's are real numbers. Then,  $Z$  has normal distribution  $N(0, \nu^2)$  with mean 0 and variance  $\nu^2 := \sum_i v_i^2$ .

We define  $\hat{X} := \frac{1}{T}(\sum_j g_j v_j)^2$  and let  $Z := \sum_j g_j v_j$ . Notice that we have  $Z \sim N(0, \nu^2)$ .

Using Fact 4.1, we can compare the moments of  $X$  and  $\hat{X}$ .

**Lemma 4.3** Define  $X$  and  $\hat{X}$  as above.

1. For all integers  $n \geq 0$ ,  $E[X^n] \leq E[\hat{X}^n]$ .
2. Using the Taylor expansion  $\exp(y) := \sum_{i=0}^{\infty} \frac{y^i}{i!}$ , we have  $E[\exp(tX)] \leq E[\exp(t\hat{X})]$ , for  $t > 0$ .

**Lemma 4.4** For  $t < \frac{T}{2\nu^2}$ ,  $E[\exp(t\hat{X})] \leq (1 - \frac{2t\nu^2}{T})^{-\frac{1}{2}}$ .

**Sketch Proof:** Observe that  $\hat{X} = \frac{1}{T}Z^2$ , where  $Z$  has normal distribution  $N(0, \nu^2)$ .

Hence, it follows that  $E[e^{t\hat{X}}] = E[\exp(\frac{t}{T} \cdot Z^2)]$ . We leave the rest of the calculation as a homework exercise.

Therefore, for  $t > 0$ , we conclude that  $E[\exp(tX)] \leq E[\exp(t\hat{X})] \leq (1 - \frac{2t\nu^2}{T})^{-\frac{1}{2}}$ , for  $t < \frac{T}{2\nu^2}$ .

**Claim 4.5** Suppose  $X := \frac{1}{T}(\sum_j \gamma_j v_j)^2$ , where  $\nu^2 = \sum_j v_j^2$ .

Then, for  $0 < t < \frac{T}{2\nu^2}$ ,  $E[\exp(tX)] \leq (1 - \frac{2t\nu^2}{T})^{-\frac{1}{2}}$ .

For negative  $t$ , we cannot argue that  $E[\exp(tX)] \leq E[\exp(t\hat{X})]$ . However, we can still obtain an upper bound using another method.

**Claim 4.6** For  $t < 0$ ,  $E[\exp(tX)] \leq 1 + \frac{t\nu^2}{T} + \frac{3}{2} \cdot (\frac{t\nu^2}{T})^2$ .

**Proof:**

We use the inequality: for  $y < 0$ ,  $e^y \leq 1 + y + \frac{y^2}{2}$ .

Hence, for  $t < 0$ ,

$$E[\exp(tX)] \leq E[1 + tX + \frac{t^2}{2}X^2] = 1 + \frac{t\nu^2}{T} + \frac{t^2}{2}E[X^2].$$

We use the fact that  $E[X] = \frac{\nu^2}{T}$ . We next obtain an upper bound for  $E[X^2]$ . From Lemma 4.3, we have  $E[X^2] \leq E[\hat{X}^2]$ .

Observe that  $\hat{X}^2 = \frac{Z^4}{T^2}$ , where  $Z$  has the normal distribution  $N(0, \nu^2)$ . Hence,  $E[\hat{X}^2] = \frac{\nu^4}{T^2}E[g^4]$ , where  $g$  has the standard normal distribution  $N(0, 1)$ .

Through a standard calculation, we have  $E[g^4] = 3$ , hence achieving the required bound. ■

## 4.2 Finding the right value for $t$ .

We now have an upper bound for  $E[e^{tX_i}]$  and hence we can finish the proof.

**Positive  $t$ .** For  $t > 0$ , we have  $Pr[Y \geq (1 + \epsilon)\nu^2] \leq \exp(-t(1 + \epsilon)\nu^2) \cdot \prod_i E[\exp(tX_i)]$   
 $\leq \exp(-t(1 + \epsilon)\nu^2) \cdot (1 - \frac{2t\nu^2}{T})^{-\frac{T}{2}}$ ,  
 where  $t$  has to satisfy  $t < \frac{T}{2\nu^2}$  too.

**Remark 4.7** In this case, the upper bound is not of the form  $E[\exp(tX_i)] \leq \exp(g_i(t))$ . Instead of trying to find the best value of  $t$  by calculus, sometimes another valid value of  $t$  is good enough.

We try  $t := \frac{T}{2\nu^2} \cdot \frac{\epsilon}{1+\epsilon}$ . In this case, we have  $(1 - \frac{2t\nu^2}{T})^{-\frac{1}{2}} \leq \sqrt{1 + \epsilon}$ . Hence,

$$Pr[Y \geq (1 + \epsilon)\nu^2] \leq (\sqrt{e^{-\epsilon}(1 + \epsilon)})^T \leq \exp(-\frac{\epsilon^2 T}{12}),$$

where the last inequality comes from the fact that for  $0 < \epsilon < 1$ ,

$$\sqrt{e^{-\epsilon}(1 + \epsilon)} = \exp(\frac{1}{2}(-\epsilon + \ln(1 + \epsilon))) \leq \exp(-\frac{\epsilon^2}{12}).$$

**Negative  $t$ .** For negative  $t$ , we use the bound  $E[e^{tX}] \leq 1 + \frac{t\nu^2}{T} + \frac{3}{2} \cdot (\frac{t\nu^2}{T})^2$ .

We can pick any negative  $t$ . So, we try  $t := -\frac{\epsilon}{2(1+\epsilon)} \cdot \frac{T}{\nu^2}$ .

$$Pr[Y \leq (1 - \epsilon)\nu^2] \leq [(1 - \frac{\epsilon}{2(1+\epsilon)} + \frac{3\epsilon^2}{8(1+\epsilon)^2}) \exp(\frac{\epsilon(1-\epsilon)}{2(1+\epsilon)})]^T.$$

We apply the inequality  $1 + x \leq e^x$ , for any real  $x$  to obtain the following upper bound.

$$[\exp(-\frac{\epsilon}{2(1+\epsilon)} + \frac{3\epsilon^2}{8(1+\epsilon)^2} + \frac{\epsilon(1-\epsilon)}{2(1+\epsilon)})]^T \leq \exp(-\frac{\epsilon^2 T}{12}).$$

One can check that  $-\frac{\epsilon}{2(1+\epsilon)} + \frac{3\epsilon^2}{8(1+\epsilon)^2} + \frac{\epsilon(1-\epsilon)}{2(1+\epsilon)} \leq -\frac{\epsilon^2}{12}$ , for  $0 < \epsilon < 1$ .

Hence, in conclusion, for  $0 < \epsilon < 1$ ,

$$Pr[|Y - \nu^2| \geq \epsilon\nu^2] \leq 2 \exp(-\frac{\epsilon^2 T}{12}). \text{ This probability is at most } \frac{1}{n^2}, \text{ if we choose } T := \left\lceil \frac{12 \ln 2n^2}{\epsilon^2} \right\rceil.$$

## 5 Lower Bound

We show that if we want to maintain the distances of  $n$  points in Euclidean space, in some cases, the number of dimension must be at least  $\Omega(\log n)$ .

### 5.1 Simple Volume Argument

Consider a set  $V = \{u_1, u_2, \dots, u_n\}$  of  $n$  points in  $n$ -dimensional Euclidean space. For instance, let  $u_i := \frac{e_i}{\sqrt{2}}$ , where  $e_i$  is the standard unit vector, where the  $i$ th position is 1 and 0 elsewhere. Then, for  $i \neq j$ ,  $\|u_i - u_j\| = 1$ .

We show the following result.

**Theorem 5.1** Let  $0 < \epsilon < 1$ . Suppose  $f : V \rightarrow \mathbb{R}^T$  such that for all  $i \neq j$ ,

$$1 \leq \|f(u_i) - f(u_j)\| \leq 1 + \epsilon.$$

Then,  $T$  is at least  $\Omega(\log n)$ .

**Remark 5.2** Observe that if we have  $1 - \epsilon \leq \|f(u_i) - f(u_j)\| \leq 1 + \epsilon$ , then we can divide the mapping by  $(1 - \epsilon)$ , i.e.  $f' := \frac{f}{1 - \epsilon}$ . Then, we have  $1 \leq \|f'(u_i) - f'(u_j)\| \leq \frac{1 + \epsilon}{1 - \epsilon} = 1 + \Theta(\epsilon)$ .

**Proof:**

For each  $i$ , consider a ball  $B(f(u_i), \frac{1}{2})$  of radius  $\frac{1}{2}$  around the center  $f(u_i)$ . Since for  $i \neq j$ ,  $\|f(u_i) - f(u_j)\| \geq 1$ , the balls are disjoint (except maybe for only 1 point of contact between two balls).

On the other hand, for all  $i > 1$ ,  $\|f(u_1) - f(u_i)\| \leq (1 + \epsilon)$ . Hence, it follows the big ball  $B(f(u_1), \frac{3}{2} + \epsilon)$  centered at  $f(u_1)$  contains all the  $n$  smaller balls.

Note that the volume of a ball with radius  $r$  in  $\mathbb{R}^T$  is proportional to  $r^T$ . Since there are  $n$  disjoint smaller balls in the big ball, the ratio of the volume of the big ball to that of a smaller ball is at least  $n$ .

Hence, we have  $n \leq \frac{(\frac{3}{2} + \epsilon)^T}{(\frac{1}{2})^T} \leq 5^T$ , for  $\epsilon < 1$ . Therefore, it follows that  $T \geq \Omega(\log n)$ . ■

## 6 Homework Preview

1. Suppose  $g$  is a random variable with normal distribution  $N(0, 1)$ . Prove the following.

- (a) For odd  $n \geq 1$ ,  $E[g^n] = 0$ .
- (b) For even  $n \geq 2$ ,  $E[g^n] \geq 1$ .

(Hint: Use induction. Let  $I_n := E[g^n] = \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} x^n e^{-\frac{x^2}{2}} dx$ . Use integration by parts to show that  $I_{n+2} = (n+1)I_n$ .)

2. Suppose  $\gamma_j$ 's are independent uniform  $\{-1, 1\}$ -random variables and  $g_j$ 's are independent random variables, each having normal distribution  $N(0, 1)$ . Suppose  $v_j$ 's are real numbers, and define  $X := (\sum_j \gamma_j v_j)^2$  and  $\hat{X} := (\sum_j g_j v_j)^2$ . Show that for all integers  $n \geq 1$ ,  $E[X^n] \leq E[\hat{X}^n]$ .
3. Suppose  $Z$  is a random variable having normal distribution  $N(0, \nu^2)$ . Compute  $E[e^{tZ^2}]$ . For what values of  $t$  is your expression valid?
4. In this question, we investigate if Johnson-Lindenstrauss Lemma can preserve area.
  - (a) Suppose the distances between three points are preserved with multiplicative error  $\epsilon$ . Is the area of the corresponding triangle also always preserved with multiplicative error  $O(\epsilon)$ , or even some constant multiplicative error?
  - (b) Suppose  $u$  and  $v$  are mutually orthogonal unit vectors. Observe that the vectors  $u$  and  $v$  together with the origin form a right-angled isosceles triangle with area  $\frac{1}{2}$ . Suppose the lengths of the triangle are distorted with multiplicative error at most  $\epsilon$ . What is the multiplicative error for the area of the triangle?
  - (c) Suppose a set  $V$  of  $n$  points are given in Euclidean space  $\mathbb{R}^n$ . Let  $0 < \epsilon < 1$ . Give a randomized algorithm that produces a low-dimensional mapping  $f : V \rightarrow \mathbb{R}^T$  such that

the areas of all triangles formed from the  $n$  points are preserved with multiplicative error  $\epsilon$ . What is the value of  $T$  for your mapping? Please give the exact number and do not use big O notation.

(Hint: If two triangles lie in the same plane (a 2-dimensional affine space) in  $\mathbb{R}^n$ , then under a linear mapping their areas have the same multiplicative error. For every triangle, add an extra point to form a right-angled isosceles triangle in the same plane.)

## References

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